



## Population attributable risk of aflatoxin-related liver cancer: Systematic review and meta-analysis<sup>☆</sup>

Yan Liu<sup>a</sup>, Chung-Chou H. Chang<sup>b,c</sup>, Gary M. Marsh<sup>c</sup>, Felicia Wu<sup>a,\*</sup>

<sup>a</sup> University of Pittsburgh, Graduate School of Public Health, Department of Environmental and Occupational Health, Pittsburgh, PA, USA

<sup>b</sup> University of Pittsburgh, School of Medicine, Department of Medicine, Pittsburgh, PA, USA

<sup>c</sup> University of Pittsburgh, Graduate School of Public Health, Department of Biostatistics, Pittsburgh, PA, USA

Available online 8 March 2012

### KEYWORDS

Aflatoxin  
Hepatocellular carcinoma  
Hepatitis B virus  
Population attributable risk  
Systematic review  
Meta-analysis

**Abstract Background:** Over 4 billion people worldwide are exposed to dietary aflatoxins, which cause liver cancer (hepatocellular carcinoma, HCC) in humans. However, the population attributable risk (PAR) of aflatoxin-related HCC remains unclear.

**Methods:** In our systematic review and meta-analysis of epidemiological studies, summary odds ratios (ORs) of aflatoxin-related HCC with 95% confidence intervals were calculated in HBV+ and HBV– individuals, as well as the general population. We calculated the PAR of aflatoxin-related HCC for each study as well as the combined studies, accounting for HBV status.

**Results:** Seventeen studies with 1680 HCC cases and 3052 controls were identified from 479 articles. All eligible studies were conducted in China, Taiwan, or sub-Saharan Africa. The PAR of aflatoxin-related HCC was estimated at 17% (14–19%) overall, and higher in HBV+ (21%) than HBV– (8.8%) populations. If the one study that contributed most to heterogeneity in the analysis is excluded, the summarised OR of HCC with 95% CI is 73.0 (36.0–148.3) from the combined effects of aflatoxin and HBV, 11.3 (6.75–18.9) from HBV only and 6.37 (3.74–10.86) from aflatoxin only. The PAR of aflatoxin-related HCC increases to 23% (21–24%). The PAR has decreased over time in certain Taiwanese and Chinese populations.

**Conclusions:** In high exposure areas, aflatoxin multiplicatively interacts with HBV to induce HCC; reducing aflatoxin exposure to non-detectable levels could reduce HCC cases in high-risk areas by about 23%. The decreasing PAR of aflatoxin-related HCC reflects the benefits of public health interventions to reduce aflatoxin and HBV.

© 2012 Elsevier Ltd. All rights reserved.

<sup>☆</sup> This work was funded by the National Cancer Institute (5R01CA153073-02) of the United States National Institutes of Health.

\* Corresponding author: Address: Department of Environmental and Occupational Health, Graduate School of Public Health, University of Pittsburgh, 100 Technology Drive, Pittsburgh, PA 15219, USA. Tel.: +1 412 624 1306; fax: +1 412 624 3040.

E-mail address: [few8@pitt.edu](mailto:few8@pitt.edu) (F. Wu).

### 1. Introduction

Aflatoxins are toxic and carcinogenic chemicals produced primarily by the fungi *Aspergillus flavus* and *A. parasiticus*, which infect food crops such as maize,

peanuts, and tree nuts. About 4.5 billion people worldwide are exposed to dietary aflatoxins.<sup>1</sup> Exposures are highest in tropical and subtropical regions of the world, where maize and peanuts are dietary staples and food storage conditions are suboptimal.<sup>1,2</sup>

Aflatoxins are amongst the most potent naturally occurring human hepatocarcinogens known. The International Agency for Research on Cancer (IARC) has classified “naturally occurring mixtures of aflatoxins” as a Group 1 human carcinogen.<sup>3</sup> Abundant epidemiological evidence suggests that aflatoxin exposure synergises with chronic hepatitis B virus (HBV) infection to increase liver cancer (hepatocellular carcinoma, HCC) risk in populations with both risk factors.<sup>4–8</sup> More recently, toxicological models for the mechanism of the synergism of these two risk factors have emerged,<sup>9–11</sup> and are summarised in Wild and Gong.<sup>12</sup> Unfortunately, both high aflatoxin exposure and HBV are prevalent in many parts of the developing world, particularly Asia and Africa.

Previously, by compiling food consumption and aflatoxin contamination data in multiple countries and conducting a quantitative cancer risk assessment, we estimated that 25,200–155,000 (5–28%) annual HCC cases worldwide could be attributed to aflatoxin exposure.<sup>13</sup> This large range highlights the limitations in obtaining exposures solely from food surveys, uncertainties in the nature of the dose–response relationship, and uncertainties in HBV prevalence data in different nations.

In this context, systematically analysing human studies that relate biomarkers of aflatoxin exposure and HBV infection to HCC may provide a more precise and accurate measurement of burden of HCC caused by aflatoxin. Therefore, in this study, we systematically reviewed epidemiological studies on these associations in different world regions. By combining the relevant odds ratios (ORs) and relative risks (RRs) from these

studies, we conducted meta-analyses to calculate population-attributable risk (PAR) of aflatoxin-related HCC in the population overall, as well as in HBV+ and HBV– populations. PAR is the proportion of disease cases that could be avoided if a particular risk factor was eliminated in a population. In the context of our study, PAR of aflatoxin-related HCC is the proportion of HCC cases that could be avoided in a chosen population by reducing aflatoxin exposures (as measured by biomarkers) from detectable to undetectable levels.

## 2. Methods

### 2.1. Search strategy

We performed a literature search until May 13th, 2011, using the following search terms on Medline/PubMed: (aflatoxin) and (hepatitis B) and (liver cancer); (aflatoxin) and (hepatitis B) and (hepatocellular carcinoma). Additionally, we searched reference lists from retrieved articles to identify further relevant studies. Our systematic review and meta-analyses were conducted in adherence to PRISMA standards for reporting meta-analyses.<sup>14</sup>

### 2.2. Eligibility criteria

Studies were included in the systematic review if they met the following criteria: (1) case-control or cohort study design; (2) aflatoxin as the exposure of interest; (3) HBV as the infection of interest (hepatitis B virus surface antigen [HBsAg] as a marker of chronic HBV infection); (4) HCC as the outcome of interest; and (5) relative risk (RR) or odds ratio (OR) estimates with 95% confidence intervals (CIs) reported, or data to calculate these.

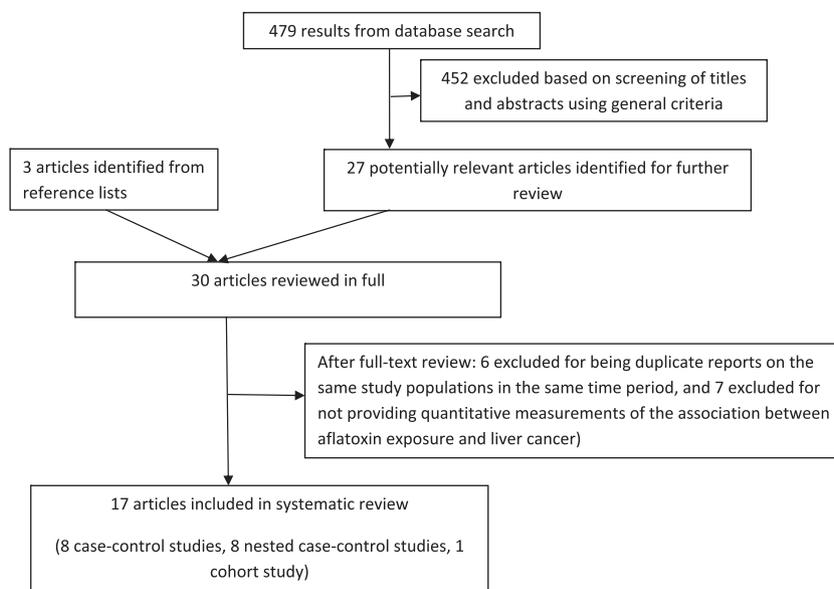


Fig. 1. Selection of studies for inclusion in systematic review.

Table 1

Characteristics of the eligible studies included in the systematic review/meta-analysis.<sup>a,b</sup>

No	Source	Location/ period	Sex	Age, yrs	No of Cases (% exposed)	No of Controls (% exposed)	Measure/Range of Exposure, detection limit	Adjusted ORs/RRs <sup>c</sup>	Adjustment for Covariates
1	Qian et al., 1994 <sup>6</sup> (cohort of 18,244 middle-aged men)	China, 1986–1992	M	45–64	50 cases (36%)	267 matched controls (12%)	AFB <sub>1</sub> -N <sup>7</sup> -Gua adduct (detectable vs non-detectable, 0.07 ng aflatoxins/ml urine)	9.1 (2.9–29.2)	HBsAg positivity, cigarette smoking
					50 cases (72%)	267 matched controls (41%)	Multiple urinary biomarker (detectable vs non-detectable, 0.01fmol/μg)	5.0 (2.1–11.8)	
2	Chen et al., 1996 <sup>18</sup> (7 township cohort nested case-control study)	Taiwan, 1991–1992	F/M	36–65	20 cases (65%)	86 matched controls (37%)	AFB <sub>1</sub> -albumin adducts (detectable Vs non-detectable, 0.01fmol/μg)	5.5 (1.2–24.5)	HBsAg, anti-HCV, family history of liver cancer cirrhosis
3	Chen et al., 1996 <sup>19</sup> (nested case-control in cohort of 4841 male HBsAg individuals)	Taiwan, 1988–1992	M	30–65	32 cases (37.5% low exposure) 32 cases (19% high exposure)	73 matched controls (33% low exposure) 73 matched controls (6.8% high exposure)	AFB <sub>1</sub> -albumin adducts (Low Vs Non-detectable, 0.01fmol/μg) AFB <sub>1</sub> -albumin adducts (High Vs Non-detectable, 0.01fmol/μg)	1.6 (0.6–4.0) 3.8 (1.0–14.5)	Cigarette smoking, alcohol consumption
4	Wang et al., 1996 <sup>7</sup> (7 township cohort nested case-control study)	Taiwan, 1991–1995	F/M	30–64	52 cases (60%)	168 matched controls (37%)	AFB <sub>1</sub> -albumin adducts (detectable vs non-detectable, 0.01fmol/μg)	1.6 (0.4–5.5)	HBsAg positivity
					38 cases (53%)	137 matched controls (45%)	Urinary aflatoxin metabolite (high vs low, 0.01fmol/μg)	3.8 (1.1–12.8)	
5	Zhang et al., 1997 <sup>22</sup> (Hospital-based case-control study)	China, 1994– 1995	F/M	18–88	152 cases (33%)	115 non-hepatic patient controls (2%)	Corn consumption history from dietary questionnaire	(1:1 pair-matched) 16.44 (1.67–61.65)	HBV infection, individual history of liver diseases, family history of liver diseases, and peanut consumption
					152 cases (89%)	115 non-hepatic patient controls (49%)	Peanut consumption history from dietary questionnaire	3.51(1.45–8.47)	
6	Yu et al., 1997 <sup>21</sup> (nested case-control of a cohort of 4841 male HBsAg individuals)	Taiwan, 1988–1994	M	30–65	42 cases (29%)	43 matched controls (14%)	AFB <sub>1</sub> -N <sup>7</sup> -gua, (below 0.21 ng/ml Vs 0.21–0.36 ng/ml, 0.05 ng aflatoxin/ml urine)	5.3(1.1–25.2)	Education level, ethnicity, habitual alcohol drinking and cigarette smoking status
					42 cases (14%)	43 matched controls (16%)	AFB <sub>1</sub> -N <sup>7</sup> -gua (below 0.21 ng/ml Vs > 0.36 ng/ml, 0.05 ng aflatoxin/ml urine)	2.8 (0.6–12.9)	
					42 cases (24%)	43 matched controls (23%)	AFM <sub>1</sub> (below 1.61 ng/ml Vs 1.61–2.85 ng/ml, 0.05 ng aflatoxin/ml urine)	1.9(0.5–7.2)	
					42 cases (55%)	43 matched controls (35%)	AFM <sub>1</sub> (below 1.61 ng/ml Vs > 2.85 ng/ml, 0.05 ng aflatoxin/ml urine)	6.0(1.2–29.0)	
7	Lunn et al., 1997 <sup>20</sup> (case-control study)	Taiwan, 1984–1995	F/M		105 cases (80%)	37 controls (43%)	AFB <sub>1</sub> -DNA adducts	Corrected OR: 3.9(1.4–11.5)	n/a
8	Kirk et al., 2000 <sup>23</sup> (case-control study)	The Gambia, 1997–1998	F/M	20–73	53 cases (36%)	53 matched controls (5.7%)	Ser-249 P53 mutation	16.4 (3.0–90.5)	Age, sex, recruitment site and HBsAg positivity
9	Sun et al., 2001 <sup>25</sup> (7 township cohort nested case-control study, HBsAg individuals)	Taiwan, 1991–1997	F/M	30–64	75 cases (64%)	140 matched controls (46%)	Aflatoxin-albumin adducts (detectable vs non-detectable, 0.01fmol/μg)	2.0 (1.1–3.7)	Sex, age and residence
10	Omer et al., 2001 <sup>24</sup> (case-control study)	Sudan 1996– 1998	F/M	20–70	115 cases	199 matched controls	Peanut butter consumption > 300 g/mo Vs Peanut butter consumption < 70 g/mo	3.3(1.4–8.1)	Age and hepatitis

(continued on next page)

Table 1 (continued)

No	Source	Location/ period	Sex	Age, yrs	No of Cases (% exposed)	No of Controls (% exposed)	Measure/Range of Exposure, detection limit	Adjusted ORs/RRs <sup>c</sup>	Adjustment for Covariates
11	Ming et al., 2002 <sup>26</sup> (Hospital-based cohort, 145 HBsAg individuals)	Qidong, China	M	27–74	31 cases	145 HBsAg + carriers follow up	AFM <sub>1</sub> (>3.6 ng/l)	3.5(1.5–8.1)	Age, HCV, family history of HCC
12	Huang et al., 2003 <sup>27</sup> (case-control study)	Qidong, China	F/M	19–87	25 cases (40%)	30 controls (6.7%)	Ser 249 TP 53 mutation	22.1(3.2–91.7)	Sex, age, recruitment site and HBsAg positivity age
13	Omer et al., 2004 <sup>8</sup> (case-control study)	Sudan, 1996– 1998	F/M	20–70	114 cases (46%)	198 matched controls (26%)	Peanut butter consumption > 300 g/mo Vs Peanut butter consumption < 70 g/mo	n/a	
14	Kirk et al., 2005 <sup>5</sup> (case- control study)	Gambia,	F/M		186 cases (40%)	348 matched controls (3.4%)	Ser-249 TP53 mutation	20.3 (8.19–50.0)	Adjusted for study group, season of recruitment and daily groundnut intake 2.11 (1.54–2.90)
15	Long et al., 2009 <sup>28</sup> (hospital-based case- control)	China, 2006– 2008	F/M		12.1% < 35, 77.8% 35–65, 10.1% > 65	618 cases (28%)	712 matched control (29%)	AFB <sub>1</sub> -adduct: Low (≤ 1.00 μmol/ mol DNA) Vs Medium (1.01– 2.00 μmol/mol DNA), (0.25 μmol/mol DNA)	
	618 cases (47%)	712 matched controls (17%)	AFB <sub>1</sub> - adduct: Low (≤		1.00 μmol/mol DNA) Vs High (≥ 2.01 μmol/mol DNA) (0.25 μmol/ mol DNA)	Age, sex, ethnicity, HBsAg, anti-HCV, and AFB <sub>1</sub> –exposure years	6.23 (4.48–8.67)		
16	Wu et al., 2009 <sup>30</sup> (7 township cohort nested case-control study)	Taiwan 1991– 2004	F/M	30–64 yr	230 cases (93%)	1052 matched controls (95%)	AFB <sub>1</sub> -albumin adduct(fmol/ mg): Non-detectable Vs Detectable (0.01fmol/μg or 1fmol/ml)	0.99 (0.48–2.02)	HBsAg, anti-HCV, habitual smoking, alcohol drinking, BMI and the batch of aflatoxin biomarker assay
					230 cases (33%)	1052 matched controls (33%)	AFB <sub>1</sub> -albumin adduct(fmol/ mg): Below the mean (<59.8) vs. Above the mean (≥ 59.8),	1.54 (1.01–2.36)	
					198 cases (88%)	904 matched controls (88%)	Urinary AFB <sub>1</sub> metabolites (fmol/ml): Non-detectable Vs Detectable (0.01fmol/μg or 1fmol/ml)	1.70 (0.89–3.25)	
					198 cases (57%)	904 matched controls (44%)	Urinary AFB <sub>1</sub> metabolites: Below the mean Vs Above the mean	1.76 (1.18–2.58)	
17	Szymanska et al., 2009 <sup>29</sup> (nested case- control study)	China, 1989– 1998	M	30–59	126 cases (67%)	123 matched controls (68%)	AF-albumin Detectable Vs non-detectable (3 pg/mg)	0.90 (0.52–1.56)	In HBV individuals

<sup>a</sup> All the eligible studies were conducted in China,<sup>6</sup> Taiwan,<sup>7</sup> or sub-Saharan Africa.<sup>4</sup> Fourteen studies reported biomarker measurements for aflatoxin exposure, while the other three studies relied on food consumption data. Twelve studies included both HBsAg+ and HBsAg– individuals, with risk estimates that were adjusted for HBsAg positivity (nine studies). Five studies were conducted in HBsAg+ populations only.

<sup>b</sup> Amongst the fourteen studies that utilised biomarkers, five measured urinary aflatoxin biomarkers, including AFM<sub>1</sub> and AFB<sub>1</sub>-N<sup>7</sup>-Guanine, six measured AFB<sub>1</sub>-albumin adducts, two measured AFB<sub>1</sub>-DNA adducts, and three measured TP53 249<sup>ser</sup> mutations. Several studies included measures of more than one biomarker.

<sup>c</sup> 15 out of 16 identified case-control studies provided matched ORs.

### 2.3. Data extraction

The following data were extracted from each study: authors, publication year, study design and sample size, study location, study period, participants' gender and age range, metric and range of aflatoxin exposure, estimated adjusted RRs/ORs and variables adjusted for analysis. Because all identified studies are case-control designs except one cohort study, and because RR and OR can be used interchangeably when the disease is relatively rare (<15%; HCC rates are lower than this in the populations studied), we combined the RR from this study with the ORs from the case-control studies to calculate a summary OR. If aflatoxin exposure was measured using different biomarkers in the same study, we selected the ones reflecting consistent biomarkers amongst different studies (one OR per study was used).

### 2.4. Statistical methods for meta-analysis

The ORs from the studies were first combined in the meta-analysis using a random-effects model, and then a fixed-effects model if heterogeneity in the study pool was insignificant.<sup>15</sup> The studies were categorised by the recruited population type: general populations, and HBV+ or HBV– populations. First, all the studies providing data for general populations (including both HBV+ and HBV– individuals) were combined, and ORs of aflatoxin-related HCC after HBsAg+ adjustment and ORs for combined (aflatoxin + HBV) effects were analysed. Then the studies with data from HBV+ populations (and studies that recruited from the general population but separately estimated ORs in HBV+ populations) were combined; and the ORs for HBV+ populations only were estimated. We also combined the studies that separately estimated the ORs in HBV– populations. If the study examined the association between aflatoxin exposure and HCC in various exposure categories, we chose the ORs reflecting highest and lowest levels of aflatoxin exposure for the meta-analysis.

Heterogeneity amongst the studies was evaluated using the Cochran's Q value calculated from the Mantel–Haenszel method and the  $I^2$  statistic.<sup>15</sup> We performed sensitivity analyses in which each study was in turn removed and the rest analysed to evaluate if the results were significantly affected by one particular study. Publication bias was assessed by a funnel plot and associated statistical tests of asymmetry. All statistical analyses were performed with Comprehensive Meta-Analysis software Version 2.2.

### 2.5. Statistical methods for PAR calculations

We estimated the PAR for aflatoxin-related HCC in HBV+ and HBV– populations for each study if the data were available. To estimate the PAR for afla-

toxin-related HCC using the adjusted ORs, we used the attributable fraction formula<sup>16</sup>:

$$AF_{POP} = \sum_{i=1}^z W_i \frac{P_i(RR_i - 1)}{1 + P_i(RR_i - 1)}$$

where  $AF_{POP}$  is aflatoxin attributable risk fraction in the population including exposed and unexposed individuals,  $P_i$  is the proportion of the population in stratum  $i$  that is exposed, and  $W_i$  is the proportion of diseased individuals (cases) in stratum  $i$ . We use adjusted  $OR_i$  in stratum  $i$  as an approximation of  $RR_i$ .

If the study provided risk estimates adjusted by HBsAg positivity, we used the formula below<sup>16</sup> to estimate the PAR of aflatoxin-related HCC in the general population:

$$AF_{POP} = \frac{P_c(RR - 1)}{RR}$$

where  $P_c$  is the proportion of cases exposed in the combined population based on detection limits for aflatoxin biomarkers in the studies, and HBsAg positivity-adjusted OR is used as an approximation of RR. For each  $AF_{POP}$ , we calculated 95% confidence intervals (CI) using the method described in Daly.<sup>17</sup>

## 3. Results

### 3.1. Literature search

The step-by-step process of our literature search is presented in Fig. 1. From 479 results, we excluded human cell line studies, animal studies, and review articles. Using the eligibility criteria described above, 27 studies were selected. Three more relevant studies were identified from the reference lists of the 27 selected studies. We then read the full texts of these 30 studies. Six studies were excluded because they were duplicated reports from the same population in the same time period, and seven more were excluded because quantitative measurements of association between aflatoxin exposure and HCC were not provided. Thus, 17 studies were included in this systematic review and PAR analysis.

### 3.2. Study characteristics

Table 1 provides an overview of the eligible studies. The 17 studies<sup>5–8,18–30</sup> on aflatoxin exposure and HCC risk – eight case-control studies, eight nested case-control studies, and one cohort study – were published between 1994 and 2009. There were 1680 HCC cases and 3052 controls in total.

Four studies reported results for one Taiwanese cohort from four different time periods<sup>7,18,25,30</sup> from 1980s to 2000s. To determine if all these studies should be included in the meta-analysis, we first examined the heterogeneity between the risk estimates provided by these studies. Because of the significant heterogeneity

Table 2  
Summary of combined odds ratios in the meta-analysis.

Risk factor	Study Population	Study area (n of studies)	Cases/controls <sup>a</sup>	Odds Ratio, 95% CI	Model	Heterogeneity
Aflatoxin only	General population with HBsAg+ adjustment	China <sup>4,6,22,27,28</sup>	634 cases/913 controls	5.99 (3.70–9.69)	Fixed	$Q = 4.86, P = 0.18, I^2 = 38.32$
		Taiwan <sup>3,7,18,30</sup> b	198 cases/904 controls	2.01 (1.40–2.89)	Fixed	$Q = 3.19, P = 0.20, I^2 = 37.29$
		Sub-Saharan Africa <sup>2,23,24</sup>	168 cases/252 controls	4.62 (2.12–10.08)	Fixed	$Q = 2.69, P = 0.1, I^2 = 62.82$
		Summary <sup>9</sup>	1000 cases/2069 controls	4.75 (2.78–8.11)	Random	$Q = 32.73, P < 0.000, I^2 = 75.56$
	General population with HBsAg+ adjustment after adjust heterogeneity	Summary <sup>8,5–8,18,22,27,28</sup>	840 cases/1302 controls	5.72 (4.42–7.40)	Fixed	$Q = 8.40, P = 0.30, I^2 = 16.66$
		Summary <sup>7,5,6,8,22,27,28,30</sup>	1000 cases/2069 controls	4.88 (2.62–9.10)	Random	$Q = 32.55, P < 0.000, I^2 = 81.57$
	General population with HBsAg+ adjustment by only including Wu et al. as follow-up for cohort in Taiwan	Summary <sup>7</sup>	1000 cases/2069 controls	4.92 (2.74–8.82)	Random	$Q = 29.48, P < 0.000, I^2 = 79.65$
	General population with HBsAg+ adjustment by taking the average effect of the series of Taiwan Studies <sup>c</sup>	Summary <sup>7</sup>	1000 cases/2069 controls	4.92 (2.74–8.82)	Random	$Q = 29.48, P < 0.000, I^2 = 79.65$
	HBsAg+ individuals	China <sup>3,6,26,29</sup>	189 cases/268 controls	2.00 (0.84–4.75)	Random	$Q = 10.66, P = 0.005, I^2 = 81.24$
		Taiwan <sup>6,7,19–21,25,30</sup>	254 cases/310 controls	1.81 (1.29–2.56)	Fixed	$Q = 8.38, P = 0.14, I^2 = 40.35$
		Sub-Saharan Africa <sup>2,5,8</sup>	128 cases/56 controls	6.48 (0.22–194)	Random	$Q = 6.54, P = 0.01, I^2 = 84.71$
		Summary <sup>11d</sup>	571 cases/634 controls	2.39 (1.50–3.82)	Random	$Q = 27.99, P = 0.002, I^2 = 64.27$
	HBsAg+ individuals after adjust heterogeneity	Summary <sup>9,5–8,19–21,25,26</sup>	377 cases/383 controls	2.90 (2.09–4.01)	Fixed	$Q = 11.16, P = 0.19, I^2 = 28.29$
	HBsAg+ individuals by only including most recent follow-up studies in a cohort of Taiwan	Summary <sup>8,5,6,8,20,21,26,29,30</sup>	571 cases/634 controls	2.27 (1.24–4.14)	Random	$Q = 24.33, P = 0.001, I^2 = 71.23$
	HBsAg+ individuals by only combing studies with adjusted ORs	Summary	332 cases/538 controls	2.10 (1.25–3.52)	Random	$Q = 16.40, P = 0.012, I^2 = 63.42$
	HBsAg+ individuals by taking the average effect of all follow-up studies in the same cohort <sup>c</sup>	Summary <sup>8</sup>	571 cases/634 controls	2.35(1.38–3.99)	Random	$Q = 23.17, P = 0.002, I^2 = 69.79$
	HBsAg– individuals	China <sup>1,6</sup>	18 cases/ 236 controls	3.4 (1.13–10.25)	/	/
Taiwan <sup>3,7,20,30</sup>		81 cases/664 controls	5.00 (2.22–11.28)	Fixed	$Q = 3.69, P = 0.16, I^2 = 45.79$	
Sub-Saharan Africa <sup>2,5,8</sup>		122 cases/391 controls	8.40 (4.15–16.99)	Fixed	$Q = 8.40, P = 0.19, I^2 = 42.63$	
Summary <sup>6</sup>		221 cases/1291 controls	5.91 (3.66–9.55)	Fixed	$Q = 7.51, P = 0.19, I^2 = 33.42$	
Summary <sup>5</sup>		172 cases/769 controls	6.37 (3.74–10.86)	Fixed	$Q = 7.11, P = 0.13, I^2 = 43.71$	
Summary <sup>6,5–8,20,30</sup>		244 cases/1072 controls	11.2 (7.48–16.7)	Fixed	$Q = 2.37, P = 0.80, I^2 = 0.00$	
HBV only	General population	Summary <sup>6,5–8,20,30</sup>	244 cases/1072 controls	11.2 (7.48–16.7)	Fixed	$Q = 2.37, P = 0.80, I^2 = 0.00$
	General population after adjusted heterogeneity	Summary <sup>5,5–8,20</sup>	171 cases/638 controls	11.3 (6.75–18.9)	Fixed	$Q = 2.36, P = 0.67, I^2 = 0.00$
Aflatoxin and HBV infection combined effects	General population	Summary <sup>6,5–8,20,30</sup>	554 cases/1456 controls	54.1 (21.3–137.7)	Random	$Q = 13.65, P = 0.02, I^2 = 63.36$
	General population after adjust heterogeneity	Summary <sup>5,5–8,20</sup>	452 cases/847 controls	73.0 (36.0–148.3)	Fixed	$Q = 3.48, P = 0.48, I^2 = 0.00$

<sup>a</sup> If there was a series of follow-up studies in the same cohort need to be combined, only the numbers of cases and controls from the largest follow-up study were counted, although different odds ratios from different follow-up studies were combined to assess the effect. All the cases and controls were only counted once, and as well as in calculations presented in Tables 4 and 5.

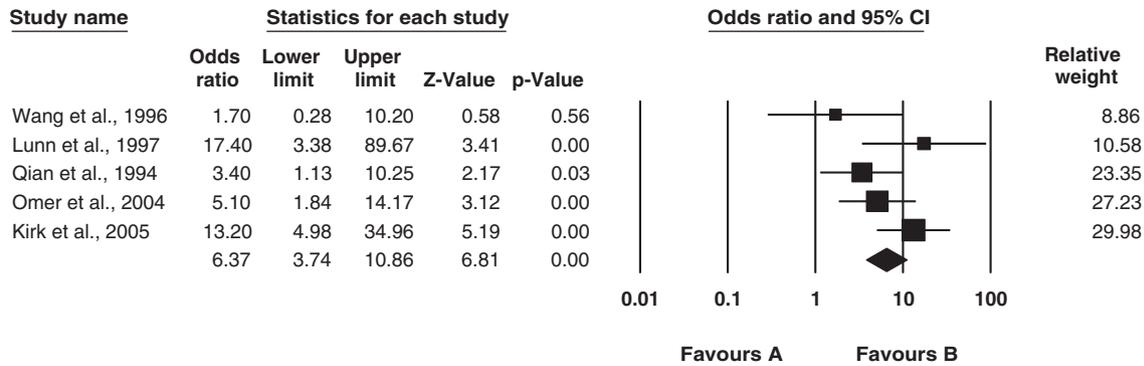
<sup>b</sup> This row shows the summary odds ratio of combing three follow-up studies in a Taiwan cohort in different years.

<sup>c</sup> The summary odds ratio obtained for the Taiwan cohort was used to represent the effect of all studies in this cohorts, and combine with other studies.

<sup>d</sup> Seven studies<sup>7,15,17,21,22,25,26</sup> reported adjusted ORs on aflatoxin-related HCC risk in HBsAg+ individuals. Four studies<sup>5,6,8,16</sup> (including two studies conducted in Sub-Saharan Africa countries) did not provide adjusted ORs directly, but provided data to calculate the unadjusted ORs. We calculated the unadjusted ORs for each of these studies and combined them with ORs from other studies with eligible data, thus we can include the effects of studies in Sub-Saharan Africa population. In subgroup analysis, the large variation of summarised ORs of aflatoxin-related HCC in HBsAg+ individuals may be explained by combining the unadjusted ORs. The heterogeneity was significant when studies were combined to examine the association between aflatoxin exposure and HCC risk in the general population and in HBsAg+ individuals.

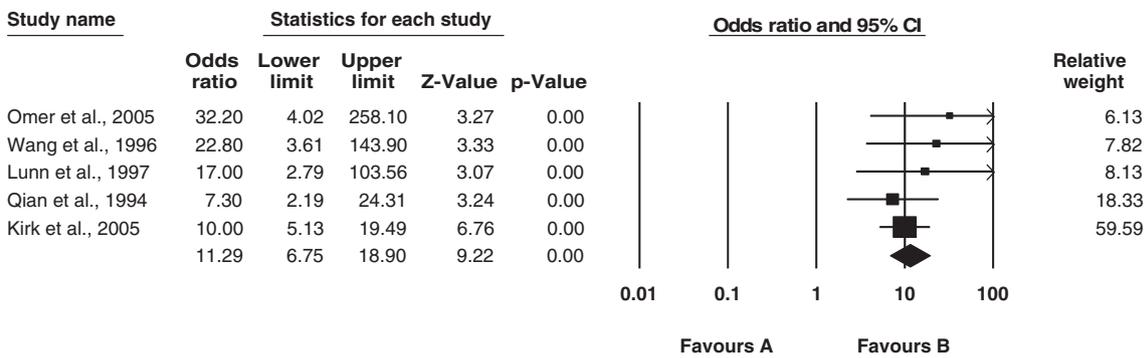
<sup>e</sup> The summary odds ratio obtained from different follow-up studies for the Taiwan cohort was used to represent the effect of all studies in this cohorts, and combine with other studies.

2A. Odds Ratios (for individual study and pooled studies) of Liver Cancer from Aflatoxin Exposure Excluding Wu et al.



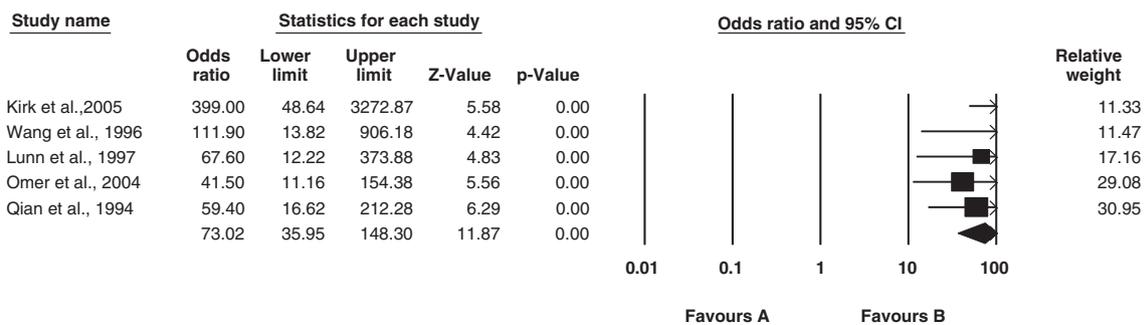
Fixed Effect Model

2B. Odds Ratios (for individual study and pooled studies) of Liver Cancer from HBV+ Excluding Wu et al.



Fixed Effect Model

2C. Odds Ratios (for individual study and pooled studies) of Liver Cancer from Combined Effects Excluding Wu et al.



Fixed Effect Model

Fig. 2. Odds ratios (ORs) and 95% CIs for association between liver cancer and two risk factors (aflatoxin exposure and chronic HBV), independently and in combination. Squares and horizontal lines correspond to the study-specific OR and 95% CI; the box size is proportional to the meta-analysis study weight; diamonds represent summarised ORs. 2A: ORs with 95% CI for association between liver cancer and chronic HBV+ only, excluding Wu et al.<sup>30</sup> 2B: ORs with 95% CI for association between liver cancer and aflatoxin exposure only, excluding Wu et al.<sup>30</sup> 2C: ORs with 95% CI for association between liver cancer and the combination effects of two risk factors, excluding Wu et al.<sup>30</sup>

of aflatoxin exposures and HCC risk estimates in this cohort between the follow-up studies through the years, we treated these as independent studies in the analysis. In analyses that included only the most recent of all studies

in a particular cohort, the results were nearly identical to those obtained when including all studies (Table 2). Two articles reported results from one case-control study in Sudan from different perspectives (risk estimates for

the general population after adjustment of HBsAg+, and risk estimates for HBsAg+ or HBsAg– separately).<sup>8,24</sup> Likewise, two articles reported results from a study in the Gambia with risk estimates for the general population after adjustment of HBsAg+, and risk estimates for HBsAg+ or HBsAg– separately.<sup>5,23</sup>

### 3.3. Aflatoxin exposure and HCC risk by HBsAg Status

The association between aflatoxin exposure and HCC, independently or in conjunction with HBV, was analysed by combining eligible studies by HBsAg+ status and calculating summary ORs (Table 2). Meta-analyses were conducted by geographic region (China, Taiwan, and sub-Saharan Africa).

Aflatoxin exposure is significantly associated with HCC risk, regardless of HBsAg status, with a summarised OR of 4.75 (2.78–8.11) from nine studies in the general population adjusted by HBsAg positivity, 2.39 (1.50–3.82) from eleven studies in HBsAg+ populations and 5.91 (3.66–9.55) from six studies in HBsAg– populations.

### 3.4. Sensitivity analysis

For the meta-analysis of aflatoxin-related HCC risk in the general population, our sensitivity analyses revealed that Wu et al.<sup>30</sup> was the most influential study in determining the summarised OR. After excluding this particular study, heterogeneity was significantly reduced ( $Q = 8.40$ ,  $P = 0.30$ ,  $I^2 = 16.66$ ), and the summarised OR was 5.57 (3.78–7.79).

For the meta-analysis of aflatoxin exposure and HCC in HBsAg+ populations, our sensitivity analyses showed that two studies, Szymanska et al.<sup>29</sup> and Wu et al.,<sup>30</sup> substantially influenced the summarised OR. After excluding the two studies, heterogeneity was significantly reduced ( $Q = 11.16$ ,  $P = 0.19$ ,  $I^2 = 28.29$ ), and the summarised OR of HCC risk for detectable vs. non-detectable aflatoxin exposure in HBsAg+ individuals was 2.90 (2.09–4.01). These results suggest that the two studies that measured the association between HCC and aflatoxin exposure in the most recent years<sup>29,30</sup> appear to have significantly different results from relatively earlier studies.

For the 10 studies<sup>6,7,18,20,22–24,27,28,30</sup> associating aflatoxin and liver cancer in the general population, we assessed publication or other forms of selection bias by a funnel plot (Fig. 2) and associated statistical tests of funnel plot asymmetry.<sup>31</sup> Seven studies are not included in this plot; five studied the association in HBsAg+ individuals only, and two are duplicate studies included in meta-analysis for different data extraction purposes, as explained in the methods. The funnel plot provides little evidence of an important departure from symmetry, indicating that publication or other forms of selection bias

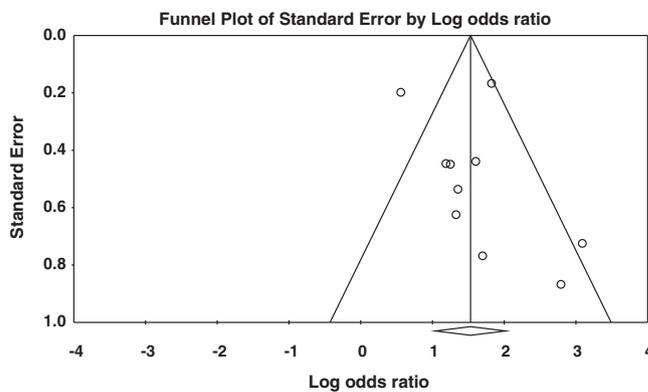


Fig. 3. Funnel plot to assess possible publication or other selection bias for the association between aflatoxin exposure and liver cancer risk in general population. No statistically significant asymmetry was found. Each circle represents 1 study. 10 studies<sup>6,7,18,20,22–24,27,28,30</sup> are eligible for this plot. 7 studies not included (5 only studied the association in HBsAg+ individuals, and 2 are duplicate studies included in meta-analysis for different data extraction purpose, as explained in the Methods section).

were not a serious limitation in our meta-analysis. This visual impression of symmetry was corroborated by the statistical tests of funnel plot asymmetry.

### 3.5. Multiplicative model of effects between aflatoxin exposure and chronic HBV infection

The meta-analysis allowed us to quantitatively evaluate the model of effects between the two risk factors aflatoxin and HBV in liver cancer. The summary OR of six studies<sup>5–8,20,30</sup> reporting ORs of HCC risk from both aflatoxin exposure and HBV is 54.1 (21.3–137.7) with significant heterogeneity ( $Q = 13.65$ ,  $P = 0.02$ ,  $I^2 = 63.36$ ). The summary OR of the same group of studies for HCC from aflatoxin exposure alone is 5.91 (3.66–9.55), while the summary OR on HCC risk from chronic HBV alone is 11.2 (7.48–16.7), both with no significant heterogeneity. When we excluded Wu et al.<sup>30</sup> which contributes most to the heterogeneity, the summarised OR for combined effects increased to 73.0 (36.0–148.3), 6.37 (3.74–10.86) for aflatoxin exposure alone, and 11.3 (6.75–18.9) for chronic HBV infection alone (Fig. 3). These estimates indicate an almost perfectly multiplicative model of effects between aflatoxin exposure and chronic HBV in HCC risk.

### 3.6. PAR of HCC from aflatoxin exposure in each study population

The PAR of aflatoxin-related HCC was calculated for each study population (Table 3). PAR is the proportion of the HCC cases that could be prevented by reducing aflatoxin exposures to “control” levels in each study. For example, HCC in the Chen et al.<sup>18</sup> Taiwanese study population could be reduced by about 10% (2.5–12%) if dietary aflatoxin exposures in this population were

Table 3

Population attributable risk of liver cancer caused by aflatoxin exposure in HBV+ populations, HBV– populations, and the general population.

Studies	Exposure measurement	PAR for aflatoxin attributable HCC risk in HBsAg+	PAR for aflatoxin attributable HCC risk in HBsAg–	PAR for aflatoxin attributable HCC risk in general study population adjusted by HBsAg+
Qian et al., 1994 <sup>6</sup> (Shanghai, China)	Multiple urinary aflatoxin metabolites	40% (24–47%) <sup>a</sup>	3.6% (0.3–5.6%)	9.0% (5.9–10.4%)
Chen et al., 1996 <sup>18</sup> (Taiwan)	AFB <sub>1</sub> albumin adducts	n/a	n/a	10% (2.5–12%)
Chen et al., 1996 <sup>19</sup> (Taiwan)	AFB <sub>1</sub> albumin adducts <sup>6</sup> Low vs undetectable	4.2% (0–13%)	n/a	n/a
	AFB <sub>1</sub> albumin adducts <sup>6</sup> High vs undetectable	4.5% (0–11%)	HBV individuals only n/a	HBV individuals only n/a
		Sum = 8.7% (0–24%)	HBV individuals only n/a	HBV individuals only n/a
Wang et al., 1996 <sup>7</sup> (Taiwan)	AFB <sub>1</sub> albumin adducts	31% (0–51%)	0 (0–2.3%)	5% (0–11%)
	Urinary aflatoxin metabolites	41% (8.1–54%)	1% (0–4.1%)	11% (1.4% - 13.7%)
Lunn et al., 1997 <sup>20</sup> (Taiwan)	AFB <sub>1</sub> -DNA adduct	31% (0–75%) <sup>b</sup>	44% (29%–47%)	n/a
Yu et al., 1997 <sup>21</sup> (Taiwan)	1.61–2.85 ng/ml AFM <sub>1</sub> vs non-detectable)	2.1% (0–7.2%)	n/a	n/a
	>2.85 ng/ml AFM <sub>1</sub> vs non-detectable)	19% (2.2–25%)	HBV individuals only n/a	HBV individuals only n/a
		Sum = 21% (2.2–32%)	HBV individuals only n/a	HBV individuals only n/a
Zhang et al., 1997 <sup>22</sup> (Henan, China)	Corn consumption	n/a	n/a	17.5% (8–18.4%)
	Peanut consumption	n/a	n/a	36% (16–45%)
Kirk et al., 2000 <sup>23</sup> (The Gambia)	Ser 249 TP53 mutation	n/a	n/a	17% (12–18%)
Omer et al., 2001 <sup>24</sup> (Sudan)	Average peanut butter consumption	n/a	n/a	23% (11–29%)
Sun et al., 2001 (Taiwan) <sup>25</sup>	AFB <sub>1</sub> albumin adducts	12% (1.7% - 20%)	n/a	n/a
			HBV individuals only n/a	HBV individuals only n/a
Ming et al., 2002 <sup>26</sup> (Qidong, China)	AFM <sub>1</sub>	57% (16–72%) <sup>c</sup>	n/a	n/a
Huang et al. 2003 <sup>27</sup> (Qidong, China)	Ser 249 TP53 mutation	n/a	n/a	17% (13–18%)
Omer et al., 2004 <sup>8</sup> (Sudan)	Average peanut butter consumption	5.4% (0–62%) <sup>d</sup>	20% (9.0–25%)	n/a
Kirk et al., 2005 <sup>5</sup> (The Gambia)	Ser 249 TP53 mutation	63% (39–67%) <sup>e</sup>	12% (6.3–17%)	13% (12–14%) <sup>6f</sup>
Wu et al., 2009 <sup>30</sup> (Taiwan)	AFB <sub>1</sub> albumin adducts	3.7% (0–11%)	1.7% (0–4.6%)	2.1% (0.06–3.4%)
	urinary aflatoxin metabolites	3.1% (0–11.7%)	4.7% (1.2–6.7%)	4.4% (1.6–6.3%)
Szymanska et al., 2009 <sup>29</sup> (Qidong, China)	AFB <sub>1</sub> albumin adducts	0 (0–14%)	n/a	n/a
			HBV individuals only n/a	HBV individuals only n/a
Long et al., 2009 <sup>28</sup> (Guangxi, China)	AFB <sub>1</sub> -DNA adduct medium vs low	n/a	n/a	6.8% (4.5–8.5%)
	AFB <sub>1</sub> -DNA adduct high vs low	n/a	n/a	19% (17–20%)
	Total	n/a	n/a	26% (22–29%)

<sup>a</sup> Calculated from unadjusted OR.<sup>b</sup> Calculated from unadjusted OR.<sup>c</sup> Author estimated.<sup>d</sup> Calculated from unadjusted OR.<sup>e</sup> Calculated from unadjusted OR.<sup>f</sup> Calculated from ORs unadjusted by HBsAg+).

reduced such that aflatoxin-albumin adduct levels were below 0.01 fmol/μg (detection limit in this study), or if dietary aflatoxin exposures could be decreased to below 4.3 ng/kg bw/day (biomarker detection limit extrapolated to dietary exposure). HCC in the study population of Shanghai males in Qian et al.<sup>6</sup> could be reduced by about 9.0% (5.9–10.4%) if aflatoxin exposures in this population were reduced to below 6 ng/kg bw/day: the

average aflatoxin exposure level in the control group. Our results showed that the PAR of HCC caused by aflatoxin is higher in HBV+ populations than in HBV– populations.

In HBV+ populations in a Taiwanese cohort, the PAR for aflatoxin-related HCC is consistently decreasing, as indicated by a series of follow-up studies: 31% in 1980s<sup>7</sup>, 12% in 1990s,<sup>25</sup> and 3% in 2000s.<sup>30</sup> Overall,

Table 4  
Estimated population attributable HCC risk from aflatoxin exposure in the general population by combining the eligible studies.

Study population		Total exposed cases (n <sub>1</sub> )	Total sample size (n <sub>2</sub> )	P <sub>c</sub> (n <sub>1</sub> /n <sub>2</sub> )	Summarised OR (95% CI)	PAR (95% CI)
General population adjusted by HBV status (6, 7, 18, 22–24, 27, 28, 30)	China	475	1588	0.299	5.99 (3.70–9.69)	25% (22–27%)
	Taiwan	113	1102	0.103	2.01 (1.40–2.89)	5.2% (2.9–6.7%)
	Sub-Saharan Africa	82	340	0.241	4.62 (2.12–10.08)	19% (13–22%)
	Summary	670	3030	0.221	4.75 (2.78–8.11)	17% (14–19%)
General population adjusted by HBV status after excluding Wu et al. 2009 <sup>30</sup>	Summary	583	2103	0.277	5.72(4.42–7.40)	23% (21–24%)

Table 5  
Estimated population attributable HCC risk from aflatoxin exposure in HBV+ and HBV– populations by combining the eligible studies.

Study population		Total HBsAg+(or HBsAg–) (n <sub>1</sub> )	Total HCC cases in HBsAg+(or HBsAg–) (n <sub>2</sub> )	Total exposed HBsAg+(or HBsAg–) (n <sub>3</sub> )	Proportion of HCC cases in HBsAg+(or HBsAg–) (W <sub>1</sub> )	Proportion of exposed HBsAg+(or HBsAg–) (P <sub>1</sub> )	Summarised OR (95%CI)	PAR (95% CI)
HBV+ population (5–8, 19–21, 25, 26, 29, 30)	China	457	189	276	0.414	0.604	2.00(0.84–4.75)	16% (0–29%)
	Taiwan	564	254	314	0.450	0.557	1.81(1.29–2.56)	14% (6.3–21%)
	Sub-Saharan Africa	184	128	76	0.696	0.413	6.48(0.22–194)	48% (0–69%)
	Summary <sup>a</sup>	1205	571	666	0.473	0.553	2.39 (1.50–3.82)	21% (10–29%)
HBV+ population after excluding Szymanska et al. 2009 <sup>29</sup> and Wu et al(30)	China	208	63	108	0.303	0.519	3.01 (1.86–4.88)	16% (9–20%)
	Taiwan	368	186	216	0.505	0.587	2.59 (1.63–4.13)	24% (14–33%)
	Summary	760	377	400	0.496	0.526	2.90 (2.09–4.01)	25% (18–30%)
HBV– population (5–8, 20, 30)	Taiwan	745	81	332	0.109	0.446	5.00 (2.22–11.28)	7% (3.8–8.9%)
	Sub-Saharan Africa	513	122	113	0.238	0.220	8.40 (4.15–16.99)	15% (9.7–19%)
	Summary	1632	227	617	0.139	0.353	5.91 (3.66–9.55)	8.8% (6.7–10%)

<sup>a</sup> Studies (including two studies in Sub-Saharan Africa countries) with unadjusted ORs were also combined to calculate the overall PAR, thus the Sub-Saharan study population can be included.

the PAR of aflatoxin-related HCC is decreasing in Taiwan in both HBV+ and HBV– individuals, from as high as 44% in 1990s<sup>20</sup> to 2% in 2000s.<sup>30</sup>

We combined all aflatoxin-exposed cases, HBV+ and HBV– individuals, and controls from all the eligible studies to calculate the PAR of aflatoxin-related HCC by HBsAg status and world region (Tables 4 and 5). The PAR of aflatoxin-related HCC in the general population after HBV adjustment is 17% (14–19%). Because the earlier sensitivity analysis demonstrated that the remaining studies after exclusion of Wu et al.<sup>30</sup> do not have statistically significant heterogeneity, we also calculated the PAR of aflatoxin-related HCC after exclusion of.<sup>30</sup> The PAR increased to 23% (21–24%).

The PAR of aflatoxin-related HCC in the HBV+ population is 21% (10–29%). A separate calculation was performed excluding Szymanska et al.<sup>29</sup> and Wu et al.,<sup>30</sup> the most influential studies indicated by the sensitivity analysis. The new PAR of aflatoxin-related HCC in the HBV+ population was 25% (18–30%). The PAR of aflatoxin-related HCC in HBV– populations is 8.8% (6.7–10%).

#### 4. Discussion

Aflatoxin exposure is significantly associated with HCC risk regardless of HBV status. Our meta-analyses show that in areas of high aflatoxin exposure and

chronic HBV infection, aflatoxin exposure and HBV have a nearly perfectly multiplicative relationship in increasing HCC risk. In populations including both HBV+ and HBV– individuals in the geographic regions studied, the PAR of aflatoxin-related HCC was estimated at 17% (14–19%). This implies that if it were possible to reduce aflatoxin to below detectable limits in these regions, HCC incidence could be reduced by 14–19%. There are roughly 520,000 new HCC cases in China, southeastern Asia and sub-Saharan Africa each year.<sup>32</sup> If the PARs are generalised to these areas, the implication is that, by reducing aflatoxin in human diets to below detectable levels, 72,800 to 98,800 new HCC cases could be prevented every year. If this PAR was generalised to regions of the world beyond Africa and Asia, the overall number of HCC cases (749,000 new cases per year<sup>32</sup>) that could be prevented by aflatoxin control would reach 105,000–142,000.

The PAR of aflatoxin-related HCC increases to 23% (21–24%), and heterogeneity amongst the studies decreases significantly, if one study<sup>30</sup> is excluded from the meta-analysis. However, this study is important because it suggests that aflatoxin exposure is decreasing over time in the Taiwanese (Penghu) population studied. Our PAR estimates for individual studies showed a decrease in PAR of aflatoxin-related HCC in the Penghu cohort in the last three decades. It is worth noting that in a 1970s food survey, over one-third of peanuts in Penghu were heavily contaminated by aflatoxins, with an average aflatoxin content of 167 µg/kg.<sup>33</sup> Mean urinary aflatoxin in HCC patients in this cohort form was 219 µg/ml in 1991/1992,<sup>7,18</sup> and decreased to 0.017 µg/ml in HCC patients in the same cohort in 2004.<sup>30</sup> Also, the HBV vaccination programme in Taiwan has successfully reduced HBV prevalence, further reducing HCC risk.<sup>34</sup>

In some parts of the world such as Taiwan, aflatoxin exposure is decreasing. In other parts of the world such as Africa, rural China, and Southeast Asia, there is little evidence that aflatoxin exposure is decreasing; in fact, two recent Kenyan events of extremely high aflatoxin levels in maize (in 2004–2005, and again in 2010) suggest the opposite. With climate change, aflatoxin contamination in food crops may become exacerbated due to the conditions favoring proliferation of *Aspergilli*.<sup>35</sup> Hence, further efforts to reduce aflatoxin-related disease are needed in high-risk areas of the world.

There are several limitations in this analysis. First, the epidemiological studies included were conducted in areas of the world with both high aflatoxin and HBV (Asia and sub-Saharan Africa). Thus, although these regions account for most of the aflatoxin-induced HCC cases worldwide,<sup>13</sup> the estimated PAR is not necessarily applicable in areas with much lower aflatoxin exposures. Second, odds ratios from studies employing food surveys, exposure biomarkers and biological effect

biomarkers were combined. This decreases the precision of the analysis, as different biomarkers have different detection limits and measure different endpoints, and food surveys are less precise than biomarkers for exposure estimation. Third, the PAR is meant to represent the proportion by which the disease could be reduced if the risk factor in question was removed. It is not possible to instantaneously reduce aflatoxin to below detectable limits worldwide – rather, the PAR calculated is meant to estimate the burden of HCC caused by one risk factor (aflatoxin) and to project the extent to which the problem could be reduced in future generations if aflatoxin control strategies were widespread.

In summary, this study is the first to quantitatively evaluate the model of effects between aflatoxin and HBV in inducing liver cancer by combining results from multiple epidemiological studies. The range of PARs calculated in this analysis, 14–19% (21–24% excluding one study contributing to heterogeneity), is consistent with our previous report of 5–28% using a different methodology (quantitative cancer risk assessment).<sup>13</sup> The PAR of aflatoxin-related HCC is higher in HBsAg+ populations than HBsAg– populations. In recent years, the PAR of aflatoxin-related HCC has shown a decreasing trend in areas such as Taiwan, indicating the benefits of reduced aflatoxin exposure and HBV prevalence by public health interventions.

### Conflict of interest statement

The authors declare that they have no competing financial interests.

### References

- Williams JH, Phillips TD, Jolly PE, et al. Human aflatoxicosis in developing countries: a review of toxicology, exposure, potential health consequences, and interventions. *Am J Clin Nutr* 2004;**80**(5):1106–22.
- Strosnider H, Azziz-Baumgartner E, Banziger M, et al. Workgroup report: Public health strategies for reducing aflatoxin exposure in developing countries. *Environ Health Perspect* 2006;**114**(12):1898–903.
- IARC. Some traditional herbal medicines, some mycotoxins, naphthalene and styrene. IARC Monographs on the Evaluation of Carcinogenic Risks to Humans 2002;**82**:171–300.
- Ross RK, Yu MC, Henderson BE, et al. Urinary aflatoxin biomarkers and risk of hepatocellular carcinoma. *The Lancet* 1992;**339**(8799):943–6.
- Kirk GD, Lesi OA, Mendy M, et al. 249ser TP53 mutation in plasma DNA, hepatitis B viral infection, and risk of hepatocellular carcinoma. *Oncogene* 2005;**24**(38):5858–67.
- Qian GS, Ross RK, Yu MC, et al. A follow-up study of urinary markers of aflatoxin exposure and liver cancer risk in Shanghai, People's Republic of China. *Cancer Epidemiol Biomark Prev* 1994;**3**(1):3–10.
- Wang LY, Hatch M, Chen CJ, et al. Aflatoxin exposure and risk of hepatocellular carcinoma in Taiwan. *Int J Cancer* 1996;**67**(5):620–5.

8. Omer RE, Kuijsten A, Kadaru AMY, et al. Population-attributable risk of dietary aflatoxins and hepatitis B virus infection with respect to hepatocellular carcinoma. *Nutr Cancer* 2004;**48**(1): 15–21.
9. Besaratinia A, Kim S-i, Hainaut P, Pfeifer GP. In vitro recapitulating of TP53 mutagenesis in hepatocellular carcinoma associated with dietary aflatoxin B1 exposure. *Gastroenterology* 2009;**137**(3):1127–37.
10. Iyer S, Groopman JD. Interaction of mutant hepatitis B X protein with p53 tumor suppressor protein affects both transcription and cell survival. *Molecular Carcinogenesis* 2011:n/a-n/a.
11. Jiang W, Wang XW, Unger T, et al. Cooperation of tumor-derived HBx mutants and p53–249ser mutant in regulating cell proliferation, anchorage-independent growth and aneuploidy in a telomerase-immortalized normal human hepatocyte-derived cell line. *Int J Cancer* 2010;**127**(5):1011–20.
12. Wild CP, Gong YY. Mycotoxins and human disease: a largely ignored global health issue. *Carcinogenesis* 2010;**31**(1):71–82.
13. Liu Y, Wu F. Global burden of aflatoxin-induced hepatocellular carcinoma: a risk assessment. *Environ Health Perspect* 2010;**118**(6): 818–24.
14. Moher D, Liberati A, Tetzlaff J, et al. Preferred reporting items for systematic reviews and meta-analyses: The PRISMA statement. *Ann Intern Med* 2009;**151**(4):264–9.
15. Petitti DB. *Meta-analysis, decision analysis, and cost-effectiveness analysis*. New York: Oxford University Press; 1999.
16. Darrow LA, Steenland NK. Confounding and bias in the attributable fraction. *Epidemiology* 2011;**22**(1):53–8.
17. Daly LE. Confidence limits made easy: interval estimation using a substitution method. *Am J Epidemiol* 1998;**147**(8):783–90.
18. Chen C, Wang L, Lu S, et al. Elevated aflatoxin exposure and increased risk of hepatocellular carcinoma. *Hepatology* 1996;**24**(1):38–42.
19. Chen CJ, Yu MW, Liaw YF, et al. Chronic hepatitis B carriers with null genotypes of glutathione S-transferase M1 and T1 polymorphisms who are exposed to aflatoxin are at increased risk of hepatocellular carcinoma. *Am J Hum Genet* 1996;**59**(1): 128–34.
20. Lunn RM, Zhang YJ, Wang LY, et al. P53 mutations, chronic hepatitis b virus infection, and aflatoxin exposure in hepatocellular carcinoma in Taiwan. *Cancer Res* 1997;**57**(16):3471–7.
21. Yu MW, Lien JP, Chiu YH, et al. Effect of aflatoxin metabolism and DNA adduct formation on hepatocellular carcinoma among chronic hepatitis B carriers in Taiwan. *J Hepatol* 1997;**27**(2): 320–30.
22. Zhang JY, Wang X, Han SG, Zhuang H. A case-control study of risk factors for hepatocellular carcinoma in Henan, China. *Am J Trop Med Hyg* 1998;**59**(6):947–51.
23. Kirk GD, Camus Randon AM, Mendy M, et al. Ser-249 p53 mutations in plasma DNA of patients with hepatocellular carcinoma from the gambia. *J Natl Cancer Inst* 2000;**92**(2): 148–53.
24. Omer R, Verhoef L, Van't Veer P, et al. Peanut butter intake, GSTM1 genotype and hepatocellular carcinoma: a case-control study in Sudan. *Cancer Causes Control* 2001;**12**(1):23–32.
25. Sun CA, Wang LY, Chen CJ, et al. Genetic polymorphisms of glutathione S-transferases M1 and T1 associated with susceptibility to aflatoxin-related hepatocarcinogenesis among chronic hepatitis B carriers: a nested case-control study in Taiwan. *Carcinogenesis* 2001;**22**(8):1289–94.
26. Ming L, Thorgeirsson SS, Gail MH, et al. Dominant role of hepatitis B virus and cofactor role of aflatoxin in hepatocarcinogenesis in Qidong, China. *Hepatology* 2002;**36**(5):1214–20.
27. Huang XHSL, Lu DD, Sun Y, et al. Codon 249 mutation in exon 7 of p53 gene in plasma DNA: maybe a new early diagnostic marker of hepatocellular carcinoma in Qidong risk area, China. *World J Gastroenterol* 2003;**9**(4):692–5.
28. Long X, Ma Y, Zhou Y, et al. XPD codon 312 and 751 polymorphisms, and AFB1 exposure, and hepatocellular carcinoma risk. *BMC Cancer* 2009;**9**(1):400.
29. Szymańska K, Chen JG, Cui Y, et al. TP53 R249S Mutations, Exposure to Aflatoxin, and Occurrence of Hepatocellular Carcinoma in a Cohort of Chronic Hepatitis B Virus Carriers from Qidong, China. *Cancer Epidemiol Biomark Prev* 2009;**18**(5): 1638–43.
30. Wu HC, Wang Q, Yang HI, et al. Aflatoxin B1 exposure, hepatitis B virus infection, and hepatocellular carcinoma in Taiwan. *Cancer Epidemiol Biomark Prev* 2009;**18**(3):846–53.
31. Borenstein F, Hedges LV, Higgins, JPT, Rothstein HR. *Introduction to Meta-Analysis*: John Wiley & Sons, Ltd; 2009.
32. Ferlay J, Shin H-R, Bray F, et al. Estimates of worldwide burden of cancer in 2008: GLOBOCAN 2008. *Int J Cancer* 2010;**127**(12): 2893–917.
33. Li KG LH, Wong SS, Chuang YS. Survey of aflatoxin contaminations in peanuts from various areas in Taiwan. Taichung: Taiwan Agricultural Chemical and Toxic Substances Research Institute. 1977:37.
34. Chen SM, Kung CM, Yang WJ, Wang HL. Efficacy of the nationwide hepatitis B infant vaccination program in Taiwan. *Journal of clinical virology: the official publication of the Pan American Society for Clinical Virology* 2011;**52**(1):11–6.
35. Wu F, Bhatnagar D, Bui-Klimke T, et al. Climate change impacts on mycotoxin risks in US Maize. *World Mycotoxin J* 2011;**4**: 79–93.